

**Complex Systems Modelling for Statistical Forecasting of  
Winter North Atlantic Atmospheric Variability: a New  
Approach**

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3 **Complex Systems Modelling for Statistical Forecasting of Winter North Atlantic**

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6 **Atmospheric Variability: a New Approach**

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## 1 **Abstract**

2 Seasonal forecasts of winter North Atlantic atmospheric variability have until recently shown  
3 little skill. Here we present a new technique for developing both linear and non-linear statistical  
4 forecasts of the winter North Atlantic Oscillation (NAO) based on complex systems modelling,  
5 which has been widely used in a range of fields, but generally not in climate research. Our  
6 polynomial NARMAX models demonstrate considerable skill in out-of-sample forecasts and  
7 their performance is superior to that of linear models, albeit with small sample sizes. Predictors  
8 can be readily identified and this has the potential to inform the next generation of dynamical  
9 models and models allow for the incorporation of non-linearities in interactions between  
10 predictors and atmospheric variability. In general there is more skill in forecasts developed  
11 over a shorter training period from 1980 compared with an equivalent forecast using training  
12 data from 1956. This latter point may relate to decreased inherent predictability in the period  
13 1955-1980, a wider range of available predictors since 1980 and/or reduced data quality in the  
14 earlier period and is consistent with previously identified decadal variability of the NAO. A  
15 number of predictors such as sea-level pressure over the Barents Sea, and a clear tropical signal  
16 are commonly selected by both linear and polynomial NARMAX models. Tropical signals are  
17 modulated by higher latitude boundary conditions. Both approaches can be extended to  
18 developing probabilistic forecasts and to other seasons and indices of atmospheric variability  
19 such as the East Atlantic pattern and jet stream metrics.

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22 **Key Words:** NAO, seasonal forecast, NARMAX, predictability, jet stream, North Atlantic,  
23 winter

## 1. Introduction

Winter North Atlantic (NA) atmospheric variability is dominated by the North Atlantic Oscillation (NAO; Hurrell and Deser, 2009; Hanna and Cropper, 2017). The NAO index gives a measure of the pressure difference between semi-permanent high pressure over the Azores and a semi-permanent low pressure over Iceland. There is a see-saw of atmospheric mass between these two nodes. The greater (smaller) the pressure difference, the more positive (negative) the NAO index. A positive NAO is associated with mild, wet and often stormy winters over northwestern Europe while a negative NAO is linked with cold, dry conditions in this region, but with wetter weather in the Mediterranean (Xoplaki et al., 2004).

The NAO can be regarded as arising from storm-track and jet-stream variability (Vallis and Gerber, 2008; Stendel et al., 2016), and is an indicator of the zonality of the atmospheric flow. A positive winter NAO tends to arise when the tropospheric jet and storm track are shifted further northwards, driving storms towards western Europe, with a more zonal jet stream, whereas a negative NAO indicates a southerly displacement with an increased meridional jet-stream component (Stendel et al., 2016) which can steer storms towards the Mediterranean and enable cold air outbreaks from the Arctic to lower latitudes. The NAO is more closely associated with shifts in jet latitude than it is with jet speed variability (Woollings et al., 2010a).

The NAO is a mode of internal atmospheric variability in idealised modelling experiments (e.g. James and James, 1989) and until recently it was considered that NAO variability on intraseasonal and interannual timescales was a result of internal atmospheric variability, or climate noise (e.g. Feldstein 2000) and largely unpredictable (e.g. Johansson, 2007; Kim et al., 2012). However, there is recent evidence that the storm track and jet stream are subject to forcing from slowly varying boundary conditions such as ocean temperatures and sea-ice

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3 55 changes, together with solar variability and influences from the winter stratosphere (for a  
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5 56 review see Hall et al., 2015).  
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10 58 A number of recent studies indicates significant potential for winter seasonal forecasting in the  
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12 59 NA region based on the influence of slowly-varying boundary conditions (Scaife et al., 2014;  
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14 60 Riddle et al., 2013; Kang et al., 2014; Dunstone et al., 2016; Stockdale et al., 2015). There are  
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16 61 also a number of older studies which identify significant skill in seasonal forecasting of the  
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18 62 winter NAO using climate models (Palmer et al., 2004; Müller et al., 2005; Derome et al.,  
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20 63 2005) and empirical approaches (Fletcher and Saunders, 2006). Scaife et al. (2014) report a  
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22 64 correlation skill of 0.62 for hindcasts of the winter NAO over the period 1993-2012, based on  
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24 65 the UK Met Office GloSea5 seasonal forecasting system (MacLachlan et al., 2015). Statistical  
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26 66 forecasts are quick and cheap to implement (e.g. Cohen et al., 2018) and therefore complement  
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28 67 the dynamical forecasts. They allow for identification of sources of potential predictability and  
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30 68 may help to explain particular instances of poor forecasts in dynamical NWP models and  
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32 69 inform their future development. Recent studies (Dunstone et al., 2016; Wang et al. 2017, Hall  
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34 70 et al., 2017) have shown promising skill in predicting the winter NAO using a linear statistical  
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36 71 framework and Folland et al. (2012) produced skillful forecasts of winter European  
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38 72 temperatures based on a number of these factors. However, these studies only use linear  
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40 73 combinations of predictors, not considering non-linear, cross-product and interaction terms.  
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42 74 Often statistical models can be constructed for a training period with a very good fit,  
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44 75 subsequently failing when making out-of-sample forecasting, due to non-stationary  
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46 76 relationships, internal variability and overfitting in the training period.  
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56 78 The slowly varying boundary conditions may act to reinforce or oppose one another and  
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58 79 numerous studies examine remote causes of NA atmospheric variability, mostly in a linear  
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3 80 framework. However, the interaction of the atmosphere with boundary forcing can be non-  
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5 81 linear (Petoukhov and Semenov, 2010), so a purely linear approach may only capture a limited  
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8 82 portion of the variability. Here we further develop statistical seasonal forecasting by using a  
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10 83 novel application of NARMAX (Non-linear Auto-Regressive Moving Average with  
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12 84 eXogenous inputs) methodology (e.g. Billings, 2013), comparing linear and polynomial  
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15 85 regression-based forecasting models. We aim to investigate whether the inclusion of non-linear  
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17 86 interactions help to explain changes in the NAO. NARMAX modelling can reveal and  
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19 87 characterize non-linear dynamic relationships among signals from recorded data and produces  
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21 88 transparent models which demonstrate how a response variable (system output signal) is linked  
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23 89 to a number of candidate explanatory variables (system input signals) and their combined  
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26 90 interactions. The NARMAX approach will construct the simplest model to explain the system:  
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28 91 therefore if a linear model provides a good representation of the system, the NARMAX method  
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30 92 will go no further (Billings, 2013, p9). NARMAX modelling was first introduced to solve non-  
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32 93 linear dynamical system identification and modelling problems in engineering, and it has been  
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35 94 successful in revealing linear and non-linear relationships at a wide range of scales within the  
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38 95 engineering, biological, ecological, medical, geophysical, and environmental sciences (e.g.  
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40 96 Billings, 2013; Bigg et al., 2014; Ayala-Solares et al., 2018).

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44 98 Here we review some of the drivers of North Atlantic climate variability identified in previous  
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47 99 research. A number of studies identify a sea-ice influence, particularly from the Barents-Kara  
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49 100 Sea (BK) region (e.g. Koenigk et al., 2016; Garcia-Serrano et al., 2017). A likely pathway of  
50  
51 101 influence is due to constructive interference of the atmospheric warming related to autumn sea-  
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53 102 ice loss with climatological planetary wave patterns (Screen et al., 2018, Wu and Smith, 2016).  
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56 103 The location of the BK region is close to the climatological ridge of the zonal wave-1 and  
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58 104 wave-2 planetary waves, with localised warming acting to reinforce this pattern (Zhang et al.,  
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3 105 2018) which enhances vertical wave propagation. This then can weaken the stratospheric polar  
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5 106 vortex with a subsequent downward propagation of this signal over a number of weeks  
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8 107 (Baldwin and Dunkerton, 2001). A stronger (weaker) vortex is associated with a positive  
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10 108 (negative) NAO, as a strengthened (weakened) circumpolar stratospheric jet induces a  
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12 109 poleward (equatorward) shift in the tropospheric jet and storm tracks (Kidston et al., 2015). A  
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14 110 cryospheric influence has also been detected from Siberian snow anomalies, which may  
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16 111 enhance the Siberian high-pressure region, resulting in vertical wave propagation into the  
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18 112 stratosphere and a weakening of the stratospheric vortex (Cohen et al., 2007). This effect has  
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20 113 been observed in models, but can be weak, and may require modulation by the Quasi-biennial  
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22 114 Oscillation (QBO) to be more effective (e.g. Tyrrell et al. 2018). The strength of the vortex can  
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24 115 also be perturbed by factors such as the QBO phase (Boer and Hamilton 2008), the solar cycle  
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26 116 (Ineson et al., 2011) and tropical volcanic eruptions (Robock and Mao, 1995; Driscoll et al.,  
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28 117 2012).

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35 119 Sea-surface temperatures (SST) are an important element of boundary layer forcing. The  
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37 120 interaction between atmosphere and ocean is complex, with the atmospheric variability  
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39 121 characterised by the NAO forcing SST variability in the NA, to produce the distinctive tripole  
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41 122 pattern of SSTs (Deser et al., 2010). However, there is evidence for feedback of this SST  
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43 123 pattern to the atmosphere at time lags of a few months, as the spring tripole anomalies are  
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45 124 preserved beneath the summer mixed layer, re-emerging in winter as the mixed layer deepens  
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47 125 (Rodwell et al., 1999, Deser et al., 2003; Czaja and Frankignoul, 1999). A complementary SST  
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49 126 pattern is the North Atlantic Horseshoe (NAH; Czaja and Frankignoul, 2002), which may  
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51 127 evolve from the tripole anomalies and where SST anomalies may lead the NAO by up to six  
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53 128 months. These patterns fluctuate on a decadal scale, but Atlantic SSTs also experience  
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3 129 multidecadal variability known as the Atlantic Multidecadal Oscillation (AMO, e.g. Enfield et  
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5 130 al., 2001) with warm and cool phases and a period of 65-80 years.  
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10 132 Other recently identified associations involve SSTs in the BK and Greenland-Norwegian (GIN)  
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12 133 Seas (Kolstad and Årthun, 2018). In addition, sea-level pressure (SLP) in the BK region can  
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14 134 precondition autumn sea-ice extent, and so itself may be a potential predictor of European  
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16 135 winter weather variability (King and Garcia-Serrano, 2016). Furthermore, the temperature  
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18 136 variability of the NA subpolar gyre (SPG) has been associated with changes in jet speed. A  
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20 137 weakened gyre circulation can lead to increased poleward transport of warmer subtropical  
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22 138 waters, and a decrease in meridional temperature gradient (Häkkinen et al., 2011; Woollings  
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24 139 et al., 2018).  
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31 141 The role of solar fluctuations in North Atlantic climate variability is a subject of considerable  
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33 142 debate. A solar cycle signal has been detected in European winters (Lockwood et al., 2010;  
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35 143 Woollings et al., 2010b), with lower solar activity associated with colder European winters.  
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37 144 Recent studies suggest that there is a lag of 3-5 years between the solar signal and its impact  
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39 145 on the atmosphere, measured by the NAO (Scaife et al., 2013; Gray et al., 2013; 2016; Andrews  
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41 146 et al., 2015), possibly as a result of integration of the solar signal over time by SSTs.  
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47 148 The influence of tropical teleconnections is also evident. The El-Niño-Southern Oscillation  
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49 149 (ENSO) signal can propagate via troposphere and stratosphere (Toniazzo and Scaife, 2006;  
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51 150 Bell et al., 2009) and there is evidence of non-linearity, with stronger El Niño events not having  
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53 151 the NAO-like impact of moderate events (Folland et al., 2012; Rao and Ren, 2016a; 2016b).  
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55 152 This may be a consequence of a more eastward centre of action for stronger events (Takahashi  
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57 153 and Dewitte, 2015) with a different pathway of propagation. In addition, links between Indian  
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3 154 Ocean SST anomalies and the NAO, (Hoerling et al., 2004; Li et al., 2010) with the NAO  
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5 155 lagging SST anomalies by a month or more and between the NAO and the Madden-Julian  
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7 156 Oscillation (Garfinkel et al., 2014 Tseng et al., 2018) have also been identified. Yu and Lin  
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9 157 (2016) find an NAO response to tropical heating anomalies in the Indian and Atlantic regions,  
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11 158 although not necessarily to SSTs.  
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17 160 Influences from these boundary conditions at a range of lead times means there could be a  
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19 161 significant component of predictability within the winter NA atmospheric circulation (Smith  
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21 162 et al., 2016). As the NAO is such a significant factor in determining the winter weather around  
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23 163 the Atlantic basin, skillful seasonal forecasts of the NAO will have considerable economic and  
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25 164 societal benefits. The NAO is related to hydrological outlooks and flood risk (e.g. Svensson et  
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27 165 al., 2015; Bell et al., 2017) and is significantly related to energy demand (Thornton et al., 2017;  
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29 166 Clark et al., 2017).  
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35 168 Section 2 presents the data used, and section 3 explains the methods, including how the  
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37 169 NARMAX methodology is applied. Results are presented in section 4 and interpreted in section  
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39 170 5. This is followed by some concluding remarks in section 6.  
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## 44 172 **2. Data**

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46 173 Data used are summarised in Table I with additional information in Table S1. HadISST1  
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48 174 (Rayner et al., 2003) is used for SST-based and sea-ice variables. We select monthly predictor  
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50 175 variables taken at lead times of one month up to a maximum of seven months (the preceding  
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52 176 May) preceding the winter in question. For this study it is assumed that there is no prediction  
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54 177 skill derived from the previous year's NAO, although November values are available for  
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56 178 selection. The sea-ice regions are taken from Screen (2017), where nine distinct sectors are  
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3 179 identified with limited covariability suggesting a large degree of regional independence. The  
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5 180 NA SST tripole index is constructed following Fan and Schneider (2012). Similarly, the NAH  
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8 181 and SST gradient indices are derived by calculating mean SSTs over two regions and  
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10 182 subtracting them. See Table S1 for details. Blanca Ayarzagüena provided the T100 index which  
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12 183 is a measure of the strength of the stratospheric polar vortex (SPV). This is an index of daily  
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14 184 temperature anomalies at 100hPa, averaged over 65-90N, derived from JRA-55 reanalysis data  
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17 185 (Kobayashi et al., 2015). Monthly mean values are constructed from an average of daily means.  
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19 186 Tropical rainfall anomalies provide an indication of convective activity and divergence aloft,  
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21 187 which can generate Rossby waves which propagate away from the source and are capable of  
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23 188 influencing extratropical atmospheric circulation (Hoskins and Karoly, 1981). Data are  
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25 189 obtained from the Global Precipitation Climatology Project (GPCPv2.3, Adler et al., 2003).  
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27 190 However, such data are not available prior to 1979 so tropical SSTs for similar regions are used  
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29 191 in the 1956 models, although there is not always a causal link between SSTs and tropical  
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31 192 rainfall, depending on the region under consideration. The MJO is another index which can  
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33 193 provide a tropical signal and is obtained from the Climate Prediction Center (CPC). Here ten  
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35 194 phases are used rather than the more common eight, and a negative value indicates the active  
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37 195 convective state (Baxter et al., 2014). QBO data (Naujokat, 1986, updated) use the 30hPa level  
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39 196 following Hamilton (1984). All data are normalised to the period 1981-2010.  
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47 198 To capture the non-linear relationship between the NAO and the ENSO signal, we use the  
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49 199 discontinuous N3.4 index of Folland et al. (2012) alongside the conventional index, where  
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51 200 values are set to zero for normalised values between  $\pm 1$ . More negative values are set to -1,  
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53 201 values in the range 1-1.75 are set to one, while values greater than 1.75 are set to zero. The  
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55 202 volcanic index used by Folland et al. (2012) is also applied. Here, for two years following a  
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57 203 volcanic eruption, values are set to one, and are zero for all other years, as Robock and Mao  
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3 204 (1995) identify the effects of eruptions in the first and second winters after the eruption. This  
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5 205 simulates the persistence of volcanic aerosols in the stratosphere after a major eruption. Major  
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7 206 tropical eruptions are as identified by Stenchikov et al. (2006).  
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12 208 Two versions of the NAO are used: the PC-based NAO (HPC; Hurrell et al., 2003) and a  
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14 209 station-based index derived from SLP data from Reykjavik and Ponta Delgada, Azores (station  
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16 210 NAO). This is taken as the difference between SLP at the two stations, which is then  
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18 211 normalised. This is the approach followed by Scaife et al. (2014), although other indices are  
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20 212 typically constructed by normalising the pressure at each station before subtracting (e.g.  
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22 213 Hurrell, 1995; Cropper et al., 2015). The correlation between these two indices for winter is  
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24 214 0.90 (1956-2017); however, this does vary slightly over time (13-year running correlations  
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26 215 between the two reach a low of 0.82 for a period in the early 1990s). The PC-based NAO better  
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28 216 captures the spatial patterns of the NAO, although is dependent on the area selected for  
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30 217 analysis, whereas due to shifts in the NAO centres of action a fixed-point index will not always  
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32 218 represent this optimally. A disadvantage of the PC-based approach is that the whole index time  
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34 219 series needs recalculating every time a new value is incorporated, and being a mathematical  
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36 220 construct, does not necessarily represent climate physics (Dommenget and Latif, 2002)  
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44 222 Data are not detrended as we aim to forecast as closely as possible to the real world. In order  
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46 223 to address any trends present in the data, indices of atmospheric carbon dioxide and global  
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48 224 temperature are available for incorporation in the models. The winter season is defined as  
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50 225 December-January-February (DJF), where the winters refer to the year of the January.  
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56 227 **3. Methods**  
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3 228 Models are constructed for training periods from 1956-2010 and 1980-2010. The year 2010 is  
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5 229 used as the cutoff, as in that year there are extreme values of the NAO and these are included  
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7 230 in the training dataset, as simulation experiments show that models trained without extreme  
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9 231 values perform poorly when used for prediction. Such extreme values are particularly  
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11 232 important in view of the relatively small sample sizes available. This leaves 2011-2018 for use  
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13 233 as a validation period (2011-2017 for the station-based NAO due to data availability). We  
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15 234 predict values for this period without adapting the model. An alternative, retroactive  
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17 235 verification approach (Mason and Baddour, 2008) constructs a model over the training period,  
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19 236 then forecasts the next year only based on that model. Forecasts for subsequent years are based  
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21 237 on models which incorporate the previous year's observation, and the models are allowed to  
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23 238 evolve, both in terms of coefficients and predictors selected. This approach was tested for linear  
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25 239 models and produced almost identical forecasts with only slight changes in coefficients, but  
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27 240 the same predictors, with no appreciable improvement in forecast quality.  
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38 242 **3.1 The NARMAX Method**  
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40 243 One of the most attractive features of the NARMAX model, distinguishing it from other non-  
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42 244 linear data-driven modelling techniques, is its power to build transparent and interpretable  
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44 245 models where the mathematical significance of each model term is meaningful (e.g. Billings,  
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46 246 2013, p10) and in most real applications the selected model terms are physically interpretable  
47  
48 247 (e.g. Billings, 2013, Chapter 14 Case Studies). Essentially, this approach treats each of the  
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50 248 candidate predictors as a possible underlying cause of change in the response variable of  
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52 249 interest and uses a set of model detection algorithms to automatically identify and pick out the  
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54 250 most important predictors, based on which it then establishes a quantitative relationship that  
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56 251 best relates the possible forcing variables to the response variable. Note that the dependence  
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252 relation of response on potential predictors can be either linear or non-linear. While traditional  
 253 linear modelling approaches such as ARMA and ARMAX might be able to capture main linear  
 254 relationships, they fail to reveal or capture any non-linear dynamics that are inherent in weather  
 255 and climate (e.g. Easterling et al., 2000; Hoerling et al., 2001; Dell et al., 2013; Burke et al.,  
 256 2015).

257

258 Taking the case of a one input ( $u$ ) and one output ( $y$ ) problem as an example, the NARMAX  
 259 model for  $y$  is written as (Billings, 2013):

$$y(k) = F(y(k-1), y(k-2), \dots, y(k-n_y), \\ u(k-d), u(k-d-1), \dots, u(k-d-n_u), \\ e(k-1), e(k-2), \dots, e(k-n_e)) + e(k) \quad (1)$$

260 where  $y(k)$  and  $u(k)$  are the measured system output (response) and input (explanatory),  
 261 respectively, at time  $k$ ;  $e(k)$  is a noise sequence which is not measurable but can be estimated  
 262 once a model is built;  $n_y$ ,  $n_u$ , and  $n_e$  are the maximum lags for the system output, input, and  
 263 noise;  $F(\bullet)$  is some non-linear function to be determined; and  $d$  is a time delay (typically  $d=0$  or  
 264  $d=1$ ). For an identified model, the noise  $e(k)$  can be estimated as the prediction errors:  
 265  $e(k) = y(k) - \hat{y}(k)$ , where  $\hat{y}(k)$  is the predicted value at time instant  $k$  generated by an  
 266 estimated model. The noise terms are included to accommodate the effects of measurement  
 267 noise, modelling errors, and/or unmeasured disturbances. For the purposes of this study, we  
 268 consider the predictability that can be obtained from the input (predictor) variables of preceding  
 269 months, and so do not consider past outputs (previous winter NAO values) further.

270

271 There are many model subset selection methods such as the traditional forward selection  
 272 (Faraway, 2002; Wilks, 2011). NARMAX uses an orthogonal forward selection algorithm,  
 273 called Forward Regression Orthogonal Least Squares (FROLS; Billings, 2013), to select the

274 important terms. The efficiency of FROLS can be attributed to its use of mutual information  
 275 (in addition to the correlation function), to measure not only linear but also potential non-linear  
 276 dependent correlation of the target signal and the candidate explanatory signals. Furthermore,  
 277 unlike traditional stepwise selection which uses hypothesis-tests and p-values to measure the  
 278 significance of variables (terms), FROLS uses an effective index called the error reduction  
 279 ratio (ERR) to measure the contribution made by each of the individual terms to explaining the  
 280 change in the response variable, based on which the most significant term will be selected in  
 281 each search step. The number of model terms can be determined using either the APRESS  
 282 (Billings and Wei, 2008) statistics or the penalized error-to-signal ratio (PESR, a variant of  
 283 APRESS) proposed in Wei et al. (2010). This is similar in principle to the Akaike Information  
 284 Criterion (AIC) and Bayesian Information Criterion (BIC) but is developed specifically for  
 285 non-linear systems. A leave-K out cross-validation is normally used with NARMAX, where K  
 286 is around 10% of the training sample. A common model is identified from subsets of the  
 287 training period, and common model parameters are then estimated using all the training data.  
 288 In this study, a leave-one-out approach is employed due to the small sample size in the training  
 289 data.

### 291 3.1.1. Linear models

292 NARMAX includes several linear model structures, e.g. autoregressive with exogenous inputs  
 293 (ARX) and autoregressive moving average with exogenous inputs (ARMAX) as a special case.  
 294 A general linear model structure of NARMAX is as follows:

$$\begin{aligned}
 y(k) = & a_1 y(k-1) + \dots + a_{n_y} y(k-n_y) + b_0 u(k-d) + b_1 u(k-d-1) \dots + b_{n_u} u(k-n_u), \\
 & + e(k) + c_1 e(k-1) + \dots + c_{n_e} e(k-n_e)
 \end{aligned} \tag{2}$$

296 where  $a_1, a_2, \dots, a_{n_y}$ ,  $b_0, b_1, \dots, b_{n_u}$ , and  $c_1, c_2, \dots, c_{n_e}$  are model parameters. The above single input  
 297 single output (SISO) case model (2) can be extended to multiple input single output (MISO)  
 298 or multiple input multiple output (MIMO) cases. The commonly used multiple linear  
 299 regression model is a special case of the MISO linear model. For example, for a case where  
 300 there are  $r$  inputs,  $u_1, u_2, \dots, u_r$ , by setting  $d=0$ ,  $n_y=0$ ,  $n_u=0$  and  $n_e=0$  yields:

$$y(k) = a_1 u_1(k) + a_2 u_2(k) + \dots + a_r u_r(k) + e(k) \quad (3)$$

### 3.1.2. Polynomial NARMAX models

303 In practice, many types of model structures are available to approximate the unknown function  
 304  $F(\bullet)$  in (1), including power-form polynomial models and rational models (Chen and Billings,  
 305 1989), radial basis function networks (Wei et al., 2007), and wavelet neural networks (Billings  
 306 and Wei, 2005; Wei et al., 2010). Power-form polynomial models are the most commonly used  
 307 representation because such models have a number of unique, attractive properties (Billings  
 308 2013, pp35-37): for example for most applications the resulting models are transparent,  
 309 physically interpretable and simple (parsimonious). This is the model form used in this study.

310  
 311 In practical applications, it is usual to consider many input signals (or explanatory variables)  
 312 and investigate how the explanatory variables (e.g.  $u_1, u_2, \dots, u_r$ ) influence the response variable  
 313 of interest. A case of many inputs can be represented by a special form of the NARMAX model  
 314 as follows:

$$y(k) = f(u_1(k), u_1(k-1), \dots, u_1(k-n_u), u_2(k), u_2(k-1), \dots, u_2(k-n_u), \dots, u_r(k), u_r(k-1), \dots, u_r(k-n_u)) + e(k) \quad (4)$$

315 where  $n_u \geq 0$ . In this study,  $n_u = 0$ , for which the model reduces to a polynomial model which  
 316 belong to the family of multiple linear regression models:

$$y(k) = f(u_1(k), u_2(k), \dots, u_r(k) + e(k) \quad (5)$$

1  
2  
3 318 For example, with two inputs  $u_1$  and  $u_2$ , the initial full model of degree 2 is:

4  
5 319 
$$y(k) = a_0 + a_1u_1(k) + a_2u_2(k) + a_3u_1^2(k) + a_4u_1(k)u_2(k) + a_5u_2^2(k) + e(k) \quad (6)$$

6  
7  
8 320 Note that in practice it may not be necessary for all the six model terms in (6) to be included in  
9  
10 321 a final predictive model, and only those that are important for explaining the variation of the  
11  
12 322 response should be included in the final model.  
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### 16 17 18 324 **3.2. Small Sample Size Problem and Model Averaging**

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21 325 The NARMAX method has been successfully applied to solve a wide range of real-world  
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23 326 problems, and in most cases the method produces a robust reliable model that can be used for  
24  
25 327 system analysis and prediction (simulation). For small sample-size problems where the number  
26  
27  
28 328 of observations is small and much smaller than the number of regressors (explanatory variables  
29  
30 329 and their cross-product interactions), the identified model can be sensitive to adding or  
31  
32 330 removing a variable or cross-product term. In order to reduce the risk of using a single model  
33  
34  
35 331 and mitigate the uncertainty in the prediction of a single model, this study proposes a simple  
36  
37 332 model-averaging approach to deal with the small sample-size problem. Models used in model  
38  
39 333 averaging only differ from each other in the number of terms used, thus all contain a common  
40  
41  
42 334 core of predictors, with only minor differences in the coefficients.  
43

44 335

45  
46 336 Assume the training dataset  $S$  contains a total of  $N$  data points. Rather than using only a single  
47  
48 337 model to calculate predictions, we use a weighted average of multiple models (in this case  
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50  
51 338 three) to carry out predictions. The model averaging scheme is described below.  
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4 340 Let  $M_1, M_2, \dots, M_s$  be the  $s$  best models identified by the PESR from the training dataset  $S$ , and  
5  
6 341 assume the values of the mean squared errors (MSE) of the  $s$  models over the training dataset  
7  
8 342 are  $mse_1, mse_2, \dots, mse_s$ , respectively. Define

10  
11 343 
$$c_1 = 1/mse_1, c_2 = 1/mse_2, \dots, c_s = 1/mse_s \quad (7a)$$

13  
14 344 
$$c = c_1 + c_2 + \dots + c_s \quad (7b)$$

15  
16  
17 345 
$$w_1 = c_1/c, w_2 = c_2/c, \dots, w_s = c_s/c \quad (7c)$$

18  
19  
20 346 Let  $\hat{y}_1, \hat{y}_2, \dots, \hat{y}_s$  be the predicted output values by the  $s$  models, the model averaging  
21  
22 347 prediction is defined as:

23  
24  
25 348 
$$\hat{y} = w_1\hat{y}_1 + w_2\hat{y}_2 + \dots + w_s\hat{y}_s \quad (8)$$

26  
27 349  
28  
29 350 **3.3. Forecast Verification**

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31  
32 351 Correlation, Mean Absolute Error (MAE) and Root Mean Square Error (RMSE) are used to  
33  
34 352 give an indication of forecast skill, for both training and testing periods. A further measure is  
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36 353 used, based on the mean squared error (MSE), the Mean Square Error Skill Score (MSESS,  
37  
38 354 e.g. Wilks, 2011, p328). This compares the skill of the forecast with a reference forecast, in  
39  
40 355 this case a forecast based on climatology. There is generally lower skill in seasonal forecasting  
41  
42 356 compared with shorter term forecasts and an out-of-sample forecast of 8 years is not large  
43  
44 357 enough to allow robust statistical conclusions to be drawn about the quality of the forecast. A  
45  
46 358 generalized discrimination score,  $D$ , (Mason, 2012; Mason and Weigel, 2009) indicates  
47  
48 359 whether the forecasts are potentially useful, despite bias and poor calibration in a low-skill  
49  
50 360 situation (Mason, 2012), based on whether forecast values increase (decrease) with an increase  
51  
52 361 (decrease) in observation values, regardless of the magnitude of the change. For a pair of  
53  
54 362 observations,  $D$  gives the probability that a forecaster can discriminate the observations based  
55  
56 363 on the corresponding forecasts.  $D$  is related to Kendall's correlation coefficient  $\tau$  by:

$$\tau = 2D - 1 \quad (9)$$

$\tau$  is effectively scaled from 0 to 1, a value of 0.5 means there is no skill in the forecast as the probability of correctly discriminating between the size of two observations is 50%. Values greater than 0.5 suggest that the forecast is potentially useful.

Comparing the forecast methods objectively and determining whether there is a significant difference is difficult with a small out-of-sample testing set ( $N=8$ ; seven for station NAO). To address this, the longer time series is split into even and odd years for the station NAO data. Models are trained on the even years ( $N=32$ ), the odd years forming the testing period ( $N=31$ ). The differences between linear and polynomial NARMAX model correlations for the testing period are assessed for significance with the “cocor” package for R (Diedenhofen and Musch, 2015), using tests for overlapping dependent samples (the correlations both concern the NAO and the compared coefficients have a shared variable, the observed NAO). The package runs six different tests and their variations, for assessing the significance of the difference between the correlations, together with a confidence interval test. For details see Diedenhofen and Musch (2015).

A further test is performed to examine whether the results obtained in the testing period for NARMAX models are likely to have occurred by chance. For the HPC80LIN model, 100 surrogate datasets (Kugiumtzis, 2000; Schrieber and Schmitz, 2000) are created for each of the predictors selected, allowing 100 surrogate NAO timeseries to be constructed for each of the models with different numbers of terms (models with five to nine terms were examined). Further details of this approach are given in the online supplementary information.

#### 4. Results

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3 389 In this section, NAO models are designated by type (station-based NAO: stat; Hurrell principal-  
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5 390 component NAO: HPC), by year, and whether linear (LIN) or polynomial (POLY) NARMAX  
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7 391 models. Thus stat80POLY is the polynomial NARMAX model for the station-based NAO,  
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9 392 using 1980-2010 as the training dataset.  
10  
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12 393

#### 14 394 **4.1 Linear Models**

16  
17 395 The predictors used in the linear models, and their coefficients are summarised in Table II and  
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19 396 verification statistics for the averaged models are shown in Table III. It is the averaged models  
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21 397 that are subsequently discussed, unless otherwise stated. An example of the model fit is shown  
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23 398 for the station-based NAO in Figure 1, with other models shown in Figure S2. There is some  
24  
25 399 consistency among the predictors selected for the different NAO indices, both within and  
26  
27 400 between the different training periods (Table II). October Barents Sea SLP and October  
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29 401 Barents-Kara Sea ice are selected for all models. Bering Sea and East Siberian-Laptev Sea ice  
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31 402 in November are selected for three of the four models; both 1980 linear models, and HPC56LIN  
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33 403 (Bering Sea ice only) and stat56LIN (East Siberian-Laptev Sea ice only). There is some  
34  
35 404 discrepancy in the cryospheric terms, with Greenland Sea ice being selected only for the 1956  
36  
37 405 models. While Greenland Sea ice is chosen in the 1956 NAO models, October Bering Sea ice  
38  
39 406 is selected for HPC80LIN as an additional term. Tropical influences are identified in both 1980  
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41 407 models, and for stat56LIN, however these show considerable variation, as variables selected  
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43 408 for the 1980 models such as the MJO and tropical rainfall, are not available for the 1956  
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45 409 models. Consequently the greater availability of tropical predictors post-1980 is reflected in  
46  
47 410 these models: November MJO (phase 8, HPC80LIN and phase 9, stat80LIN) and tropical  
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49 411 Atlantic rainfall (July; stat80LIN) and August (HPC80LIN) are present in both 1980 linear  
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51 412 models. Despite being previously identified as an important predictor of winter extratropical  
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53 413 atmospheric variability, N3.4 is not particularly prominent, only being selected for both station-  
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3 414 based NAO models (July N3.4I). Extratropical oceanic SST-related terms are selected for all  
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5 415 except stat80LIN; October GIN SST for HPC80LIN and NAH (September-stat56LIN,  
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7 416 October-HPC56LIN). The only model to detect a stratospheric influence is HPC56LIN, where  
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10 417 the stratospheric polar vortex indices for October and November are selected, although  
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12 418 interestingly these are of opposite signs and are only weakly correlated ( $r=0.21$ ).

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17 420 Correlations of the 1980 models for the 2011-2018 testing period (Table IIIa) are 0.90  
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19 421 (HPC80LIN) and 0.82 (stat80LIN). However, correlations can mask a systematic negative bias  
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21 422 in the forecasts, particularly in the linear model forecasts (with the exception of HPC80LIN),  
22  
23 423 which becomes evident when MAE and RMSE are examined. For example, although  
24  
25 424 stat80LIN correlates very well in the testing period ( $r=0.82$ ), MAE and RMSE are 1.11 and  
26  
27 425 1.26 respectively, considerably higher than the training period values. MESS scores are  
28  
29 426 negative, the forecasts having less skill than a climatological forecast. The correlations indicate  
30  
31 427 that the two NAO models are able to capture aspects of the local maxima and minima of the  
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33 428 observed NAO during the testing period, but with a large negative bias (Figure 1), which  
34  
35 429 reduces skill according to MESS and inflates the MAE and RMSE. However, D-scores  
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37 430 suggest that the models for the NAO have considerable potential usefulness (0.80, stat80LIN;  
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39 431 0.75, HPC80LIN, Table IIIa).

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45 433 For the 1956 linear models, over the testing period the correlation is weaker for both  
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47 434 (HPC56LIN, 0.46; stat56LIN, 0.69; Table III) compared with the 1980 models although no  
48  
49 435 significant difference can be determined due to small sample size. MAE and RMSE values for  
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51 436 test data are similar to the 1980 linear model for the station NAO, but greater for the HPC  
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53 437 NAO, but are always larger than those for the training period, reflecting the negative bias of  
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3 438 forecasts in the testing period. MESS are negative, but D-scores indicate potential usefulness  
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5 439 (0.80, stat56LIN; 0.68, HPC56LIN).  
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9  
10 441 Many of the terms used in construction of the linear models, for example sea-ice, SST and  
11  
12 442 tropical rainfall terms contain significant trends, which are likely to contribute to the negative  
13  
14 443 bias of the forecasts. It is notable that the station NAO forecasts from 1980 training data appear  
15  
16 444 to have overall negative trends (Figure 1), probably due to the influence of the sea-ice inputs  
17  
18 445 to the models, as found in Hall et al. (2017).  
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21 446

## 22 447 **4.2 Polynomial Models**

23  
24 448 The polynomial models are presented in detail in Table IV and compared with linear models  
25  
26 449 in Figures 1 and S2. The comparisons use the averaged linear and polynomial models and an  
27  
28 450 example of model averaging for the stat80POLY model is shown in Figure S3. Verification  
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30 451 statistics are in Table III. For the HPC80 forecast the NARMAX algorithm produces a linear  
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32 452 model only, as the inclusion of interaction terms does not significantly improve prediction  
33  
34 453 performance.  
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40 455 There is less consistency amongst the predictor variables selected (Table IV) compared with  
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42 456 the linear models, as terms are more complex and often based on interactions between  
43  
44 457 variables, and with short datasets the models are very sensitive to slight differences in  
45  
46 458 predictors. Equally, different predictors selected may be trying to capture the same variability  
47  
48 459 or acting as proxies for some hidden variable. Some predictors are selected for both linear and  
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50 460 polynomial models but may differ by a month or two.  
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3 462 Forecasting verification statistics for the polynomial model averages are shown in Table III. In  
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5 463 almost all cases, MAE, RMSE, MSESS and correlation are superior to the linear models in  
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7 464 both training and testing periods, for both time periods. The D-score is consistently greater than  
8  
9 465 0.7, indicating considerable potential usefulness of the forecasts and MSESS are always  
10  
11 466 positive. Verification statistics are mostly poorer for the 1956 models. Initially this may appear  
12  
13 467 to be surprising as the models are constructed over a longer training set, although data from the  
14  
15 468 earlier part of the time series prior to the satellite era are of poorer quality, and some variables  
16  
17 469 are not available for the earlier period. See also section 5.1 on variable inherent predictability  
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19 470 over the time period. The 1956 polynomial models show considerable improvements in all  
20  
21 471 verification metrics compared with linear models on the longer time-scale, although results  
22  
23 472 from linear models are broadly comparable to results obtained using the shorter training set.  
24  
25 473 Due to the small sample size, the statistical significance of any improvements achieved by  
26  
27 474 polynomial over linear models cannot always be satisfactorily determined. As would be  
28  
29 475 expected for such a small sample size, when correlations in the testing period are compared  
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31 476 using the suite of tests in the “cocor” package, differences are not significant.  
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40 478 In order to create a longer testing period and provide greater confidence in the performance of  
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42 479 the models in out-of-sample prediction, the 1956-2017 period was split into even years for the  
43  
44 480 training dataset and odd years for the testing dataset, using the station-based NAO index. The  
45  
46 481 predictive skill of NARMAX polynomial and linear models when even years only are used as  
47  
48 482 the training period is shown in Figure S4 and verification statistics are presented in Table S2.  
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50 483 Correlations for the testing period (odd years) are 0.45 (linear) and 0.60 (polynomial). While  
51  
52 484 both are significant ( $p < 0.05$ ), the polynomial correlation coefficient is appreciably higher,  
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54 485 consistent with results from the shorter testing datasets, yet still not deemed significantly so by  
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56 486 the range of comparison tests in “cocor”. However, this does not tell the whole story. For the  
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3 487 earlier part of the testing period (1957-1979), correlations for the linear and polynomial models  
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5 488 are very similar (0.6, linear; 0.66, polynomial), however in the later period (1981-2017), the  
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7 489 correlations are significantly different using “cocor” (0.04, linear; 0.44 polynomial,  $p < 0.1$ , one-  
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9  
10 490 tailed test). In addition both MAE and RMSE for 1981-2017 are around two thirds of the values  
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12 491 for 1957-1979. This is suggestive of some change in the NAO in the later period, which is  
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14 492 better captured by the polynomial models.  
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19 494 Verification statistics taken over the 100 surrogate models for HPC80LIN show that model  
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21 495 performance is much poorer for the surrogates in the testing period (Table S3), with significant  
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23 496 differences between model performance and the surrogate data for all verification metrics. This  
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25 497 indicates that the good prediction performance obtained with NARMAX is not a result of  
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27 498 chance but is due to the efficacy of the algorithm.  
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33 500 Figure 2 allows closer inspection of model performance in the testing period. Both linear and  
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35 501 polynomial 1980 models capture the local maxima and minima of both versions of the NAO  
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37 502 for the first three years (Figure 2a,b), although the amplitude of the response is much reduced  
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39 503 in the stat80LIN model (Figure 2b), which also display a negative bias, as discussed above,  
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41 504 while any bias in stat80POLY is minimal and non-systematic. Similarly 1956 linear models for  
42  
43 505 both NAO indices reproduce the interannual change from 2011-2014, after which the response  
44  
45 506 is damped. In general, polynomial models better match the local maxima and minima, with  
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47 507 the exception of HPC56POLY which forecasts a negative rather than positive NAO in 2014  
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49 508 (Figure 2c).  
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56 510 Figures 3 and 4 give a visual representation of the selected predictors for the different NAO  
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58 511 models, for both training periods, in order to highlight common predictor variables, although  
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3 512 some of these are used in interaction terms in the NARMAX models. For 1980 (Figure 3),  
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5 513 Barents Sea SLP and Bering Sea ice are used in all three models, Barents-Kara Sea ice, East  
6  
7 514 Siberian-Laptev Sea ice and tropical Atlantic rainfall are selected by both linear models while  
8  
9 515 MJO phase 8 occurs for both linear and polynomial HPC models. For the 1956 models (Figure  
10  
11 516 4), some predictors (Greenland Sea ice, NAH, Barents Sea SLP) are included in all four models,  
12  
13 517 while October Barents-Kara Sea ice and the N3.4I index are selected in three. A number of  
14  
15 518 predictors occur in only one model, due to them occurring as interaction terms, and the  
16  
17 519 sensitivity of the model to slight changes in variables given the short time series. For the 1956  
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19 520 training data, linear and polynomial models will select predictors that were suboptimal in the  
20  
21 521 1980 models, as the best predictors are not always available for the longer time series. Thus  
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23 522 NARMAX selects N3.4I for the 1956 models, whereas for the 1980 models the optimal tropical  
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25 523 influences seems to come from the MJO and tropical rainfall. While models are capturing  
26  
27 524 essentially the same signals, there is high sensitivity to small variations in input data due to  
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29 525 small sample size: hence different predictors are selected, which represent slightly different  
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31 526 aspects of a common signal.  
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40 528 From the results, it is remarkable that while linear models show more limited skill, particularly  
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42 529 over the longer training dataset, polynomial model forecasts, whether linear or non-linear,  
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44 530 particularly those based on the 1980 training set, are able to replicate local maxima and minima,  
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46 531 the amplitude of the observation and have minimal bias for the testing dataset.  
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## 51 533 **5. Discussion**

### 52 534 **5.1 Differences when using longer and shorter training periods**

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55 535 Based on the verification statistics (Table III), polynomial models generally outperform the  
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57 536 equivalent linear model (lower MAE and RMSE; higher correlation, MSESS and D-score),  
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3 537 although the significance of this is hard to assess on a small sample size. 1956 models generally  
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5 538 perform less well than the equivalent from 1980. There is evidence from other studies to  
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7 539 suggest that from the 1950s to 1960s there is less skill in dynamical forecasting hindcasts, with  
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9 540 the possibility that during this period the atmospheric circulation was inherently less  
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11 541 predictable (e.g. Weisheimer et al., 2016), associated with a more prevalent negative NAO.  
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13 542 However, dynamical models appear to be skillful in predicting strongly negative NAO events,  
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15 543 but less skillful regarding weaker negative NAO events. O'Reilly et al. (2017) concur that in  
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17 544 the mid 20<sup>th</sup> century, forecast skill for the NAO is reduced, and attribute this to weaker forcing  
18  
19 545 from tropical Pacific SSTs during this period. An alternative perspective is presented by  
20  
21 546 Woollings et al. (2015). Decadal variability of the NAO is identified, which is associated with  
22  
23 547 changes in strength of the jet and different dynamical behaviour (eddy-mean flow interaction,  
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25 548 Rossby wavebreaking and blocking), the period from 1980 being associated with relatively  
26  
27 549 higher jet speeds. Therefore when statistical models are trained on a longer period, with the  
28  
29 550 assumption of stationarity, these decadal variations are averaged out and predictors that are  
30  
31 551 selected are likely to be sub-optimal. Furthermore, if there is a return to relatively weak forcing  
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33 552 from Pacific SSTs, or a period of reduced jet speeds, forecasts trained on the recent period may  
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35 553 perform less well (O'Reilly et al., 2017). Weisheimer et al. (2016) report that forecast skill  
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37 554 actually increases further back in the 20<sup>th</sup> century, in the 1930s and 1940s, suggesting the issue  
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39 555 may not be attributed to reduced data quality. This corresponds to the decadal fluctuations  
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41 556 identified by Woollings et al. (2015), suggesting that periods of lower jet speed may contribute  
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43 557 to the reduced predictability. In this study, a longer training dataset does not equate to a better  
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45 558 forecasting model, in agreement with the studies above. It is also notable that when tested on  
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47 559 all the odd years, a polynomial model significantly outperforms a linear model for the period  
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49 560 1981-2017, better capturing the NAO for both early and late periods (Figure S4, Table S2).  
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3 562 Greenland Sea ice is a predictor in all 1956 models (Figure 4) but is not selected for models  
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5 563 based on the shorter training period. The months selected (May-July) represent the time of  
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7 564 maximum decrease in the regional annual sea-ice cycle (October is selected as the 8<sup>th</sup> and final  
8  
9 565 term in stat56LIN, Table IIc, and makes a minimal contribution). Interannual variability and  
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11 566 mean values of Greenland Sea ice are notably larger prior to 1980; consequently contributions  
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13 567 of the predictor terms to the models are greatest prior to 1980. Given the relatively small  
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15 568 coefficients used in the models, contributions between 1980 and 2010 are negligible, also  
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17 569 indicating why this variable is not selected for the 1980 models. Figure S5 shows that the four  
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19 570 Greenland Sea ice terms used essentially make similar contributions to each model, and can be  
20  
21 571 regarded as different, somewhat imperfect representations of some unknown predictor variable.  
22  
23 572 It is possible that the association in the 1960s and 1970s is related to the Great Salinity Anomaly  
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25 573 (e.g. Dickson et al., 1988), which circulated in the North Atlantic at this time and had its origins  
26  
27 574 in increased ice export and freshwater release through the Greenland Sea region. The  
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29 575 association between the Greenland Sea ice and NAO is physically plausible at this time,  
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31 576 particularly as the predictors capture the transition from a more negative to a positive NAO  
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33 577 over the period.  
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## 42 579 **5.2. Tropical forcing and interaction terms**

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44 580 Evidence for forcing of the winter NAO from the tropics is evident in all models; the 1980  
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46 581 training data favours the selection of the MJO and tropical rainfall, whereas these are not  
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48 582 available for the 1956 training data, so the N3.4 Index becomes more prominent, along with  
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50 583 West Indian Ocean SST in stat56LIN and HPC56POLY. It is also notable that HPC56LIN  
51  
52 584 includes no tropical forcing, although the use of stratospheric polar vortex terms here could  
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54 585 incorporate aspects of tropical forcing, such as the ENSO-stratospheric-mid-latitude  
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56 586 teleconnection (e.g. Bell et al., 2009). Tropical variables selected are somewhat inconsistent,  
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3 587 in terms of both month and predictor, which is likely to be a result of model sensitivity to slight  
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5 588 variations due to the short time series used. However, the evidence for tropical signals is clear.  
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10 590 There are multiple interaction terms identified in the polynomial models, which can be very  
11  
12 591 difficult to explain effectively without subsequent further analysis. N3.4I terms are selected in  
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14 592 both 1956 polynomial models, possibly due to the unavailability of MJO indices as discussed  
15  
16 593 above. However, in each case it occurs as an interaction term with a high latitude predictor  
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18 594 (terms 9 and 10, stat56POLY; terms 4 and 8, HPC56POLY, Table IVb,c). A closer inspection  
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20 595 of the interaction terms' time series reveals that the main input parameter is N3.4I, with the  
21  
22 596 high latitude term modulating the signal in terms of amplitude and sometimes polarity (not  
23  
24 597 shown). This, along with term 1 in stat56POLY (Table IVb), is suggestive of a tropical signal,  
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26 598 propagated by Rossby waves, which is then modulated by slowly-varying boundary conditions  
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28 599 at higher latitudes (e.g. Ding et al., 2014). Term 5 in HPC56POLY (June Hudson Bay ice x  
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30 600 volcanic index; Table IVc) operates in a similar way, the magnitude of the tropical volcanic  
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32 601 signal being the dominant input, modulated by the sea-ice term.  
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41 603 A composite plot of low minus high years for MJO phase 8 in October shows a wave train  
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43 604 emanating from the central Pacific, over North America and to the Atlantic (Figure S6).  
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45 605 Poleward propagating Rossby waves emanate from only a few source regions, whose activity  
46  
47 606 varies with tropical rainfall variability from year to year (Scaife et al., 2017). Of note in this  
48  
49 607 wavetrain is a node boundary near the Labrador Sea. An interaction term between September  
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51 608 Labrador Sea ice and the October MJO8 is identified for stat80POLY (Figure 3, Table IVa). It  
52  
53 609 is possible that in years with high sea-ice in this region, there is an interaction with the MJO  
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55 610 signal, and the effect of this interaction is sustained into winter. As the interaction term is  
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57 611 multiplicative, then it is the sign of each index which is particularly important. Both high (low)  
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3 612 sea-ice and high (low) MJO8 will combine to make the NAO prediction more positive. If,  
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5 613 however, one input is positive and the other negative, that will combine towards a more  
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7 614 negative NAO forecast.  
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### 12 616 **5.3. Higher latitude forcing and interaction terms.**

14 617 Both cryospheric and extratropical forcings are included in all models, often as interaction  
15 618 terms with tropical forcing as discussed above. Key sea-ice predictors are October Barents-  
16 619 Kara Sea (five out of seven models), November East Siberian -Laptev Sea (four out of seven),  
17 620 Bering Sea (five out of seven). The Greenland Sea has been discussed in Section 5.1 above.  
18 621 Sea-ice in the Barents-Kara Seas has frequently been identified as a key source of potential  
19 622 winter NAO predictability (Scaife et al., 2014; Garcia-Serrano et al., 2015; Wang et al., 2017;  
20 623 Hall et al., 2017). The importance of Bering Sea ice is more surprising, but could well be a  
21 624 proxy for atmospheric variability in the region, such as the Pacific-North American pattern  
22 625 (PNA), with associations with North Atlantic atmospheric variability (e.g. O'Reilly et al.,  
23 626 2017), which in turn is linked to tropical Pacific SST variability.  
24 627

25 628 Another input variable worthy of more detailed discussion is Barents SLP. This is selected by  
26 629 all NAO models (Figures 3, 4). The most commonly selected month is October (four times).  
27 630 In polynomial models Barents Sea SLP occurs as an interaction term with tropical signals  
28 631 (stat80POLY, Table Iva, HPC56POLY, Table IVc), as a single term (stat56POLY, Table IVb),  
29 632 an interaction term with SSTs (stat56POLY, Table IVb) and the QBO (HPC56POLY, Table  
30 633 IVd). A cyclonic (anticyclonic) anomaly in the BK region in October can lead to positive  
31 634 (negative) sea-ice anomalies there in November (King and Garcia-Serrano, 2016). In other  
32 635 words, the preceding pressure/geopotential height anomaly is a precursor of the Barents-Kara  
33 636 Sea ice ice which has been frequently identified as a predictor in other studies (e.g. Garcia-  
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3 637 Serrano et al., 2015; Hall et al., 2017). However, it also has the advantage that it is not subject  
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5 638 to the same long-term dramatic trends as autumn sea-ice. The interaction terms support the  
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7 639 concept of geopotential height anomalies over the BK Seas in autumn modulating signals from  
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9 640 the tropics, stratosphere and cryosphere in a non-linear way (e.g. Vihma et al., in review,  
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11 641 International Journal of Climatology). There are a number of input terms from these high  
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13 642 latitude regions that are combined in polynomial models: [November east Siberian Sea ice x  
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15 643 November GIN seas SST], [October Barents SLP x November GIN seas SST] (stat56POLY);  
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17 644 [September GIN seas SST x October Atlantic SST gradient] (HPC56POLY). It seems likely  
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19 645 that these terms are capturing different aspects of autumn variability in the northern seas, that  
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21 646 would merit further investigation as important predictors of the winter NAO.  
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28 648 Labrador Sea ice is only selected as a predictor in polynomial models, as an interaction term  
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30 649 with tropical (see section 5.2 above) or extratropical forcings. An interaction term selected in  
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32 650 both HPC56POLY and stat80POLY is [October Labrador Sea ice x October NAH SST  
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34 651 pattern]. The NAH pattern of SSTs is associated with forcing of the winter NAO by persistent  
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36 652 SST anomalies at up to six months lead time (Czaja and Frankignoul, 2002). The separate input  
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38 653 terms and the resulting multiplicative term are shown in Figure S7a. An examination of the 13-  
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40 654 year running correlation between the October NAH and winter NAO reveals a consistent  
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42 655 negative correlation, except for the period from 1995 to 2005 when the correlation coefficient  
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44 656 increases sharply, becoming positive, before a rapid return to negative values (Figure S7b).  
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46 657 This positive excursion coincides with large positive Labrador Sea ice anomalies in October,  
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48 658 these being negative for the rest of the time series. The regions for both these variables have a  
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50 659 partial overlap. Therefore when there is a negative ice anomaly in the Labrador Sea, the  
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52 660 inverted NAH index provides a predictor of the winter NAO interannual variability; however  
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54 661 for the brief period with positive ice anomalies, there is a positive relationship between the  
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3 662 NAH and NAO interannual variability (Figure S7a). The sea-ice in the Labrador Sea modulates  
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5 663 the NAH/NAO interaction, or there is a hidden North Atlantic variable for which Labrador Sea  
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8 664 ice is a proxy.  
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## 11 666 **6. Summary**

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14 667 The NARMAX approach shows appreciable potential skill in out-of-sample forecasting, albeit  
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16 668 with small testing datasets, for both linear and polynomial models which both outperform a  
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18 669 more conventional ordinary least squares approach to multiple regression (e.g. Hall et al.,  
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20 670 2017). There are strong correlations with observations, reproducing local maxima and minima  
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22 671 of the observations, and the amplitude of the observed signal. Model fits are improved when  
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24 672 based on a shorter training dataset from 1980-2010. This may partly relate to a wider range of  
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26 673 potential predictors being available for this period, but is also because of reduced inherent  
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28 674 predictability of circulation indices during the 1950s and 1960s. The skill of polynomial models  
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30 675 is consistently greater than that of equivalent linear models, and error statistics are reduced, but  
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32 676 small sample size means that further work is needed to confirm the significance of this result.  
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34 677 However, an analysis based on using odd years for the testing data is strongly suggestive of a  
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36 678 better performance by a polynomial model over that of a linear model, particularly in that it  
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38 679 better represents the transition from the early more negative NAO period to the end of the  
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40 680 1970s to the more positive phase post-1980. NARMAX can identify important predictors of  
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42 681 winter North Atlantic atmospheric variability. Discrepancies between predictor selection in  
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44 682 models is likely to arise through increased sensitivity to small fluctuations in input, due to the  
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46 683 small sample sizes available. This means that the models capture the same signals, but select  
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48 684 them in slightly different ways.  
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3 686 An important result of the study is that polynomial NARMAX models are capable of revealing  
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5 687 the potential modulation of tropical forcing by higher latitude boundary conditions. Barents  
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7 688 Sea SLP can also play a crucial role in modulating cryospheric and extratropical signals in  
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9 689 addition to those from the tropics. This could be significant in developing the next generation  
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11 690 of NWP models.  
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17 692 Our NARMAX approach can be extended to other circulation indices such as the East Atlantic  
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19 693 and Scandinavian Patterns and jet latitude and speed, and to other seasons. This may be  
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21 694 especially beneficial for summer seasonal forecasts, where there is currently relatively little  
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23 695 predictability from dynamical models (e.g. Ossó et al., 2018; Dunstone et al., 2018). It is also  
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25 696 possible to extend the approach to probabilistic forecasting and - by utilising links between  
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27 697 North Atlantic circulation patterns and, for example, UK regional temperature and precipitation  
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29 698 patterns (Hall and Hanna, 2018) - provide enhanced seasonal forecasts that should be useful  
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31 699 for a wide range of stakeholders. NARMAX can also be used to assess how contributions from  
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33 700 different atmospheric circulation predictors vary over time using a moving window approach.  
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35 701 Future work will use Coupled Model Intercomparison Project (CMIP) 5/6 output to construct  
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37 702 models using a longer timeseries, enabling the use of a longer testing dataset, to confirm  
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39 703 whether polynomial forecasts are significantly improved compared to linear versions. The  
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41 704 models can be further extended to include previous years' predictor values, and by increasing  
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43 705 the lead-time at which forecasts are issued for a given season.  
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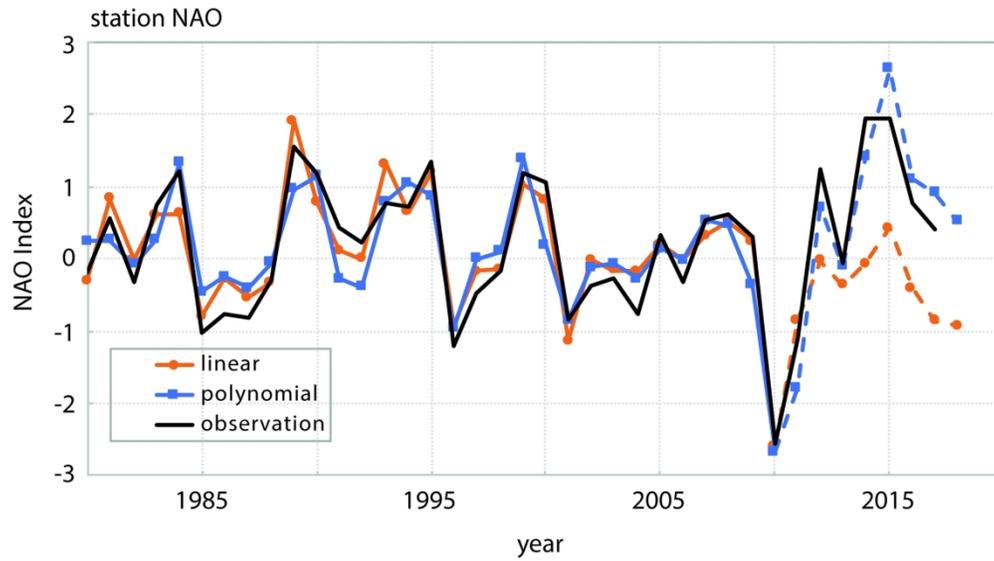


Figure 1 Example of hindcasts (solid) and forecasts (dashed) using linear (LIN) and polynomial (POLY) NARMAX models, derived from the 1980 training period for the station-based NAO.

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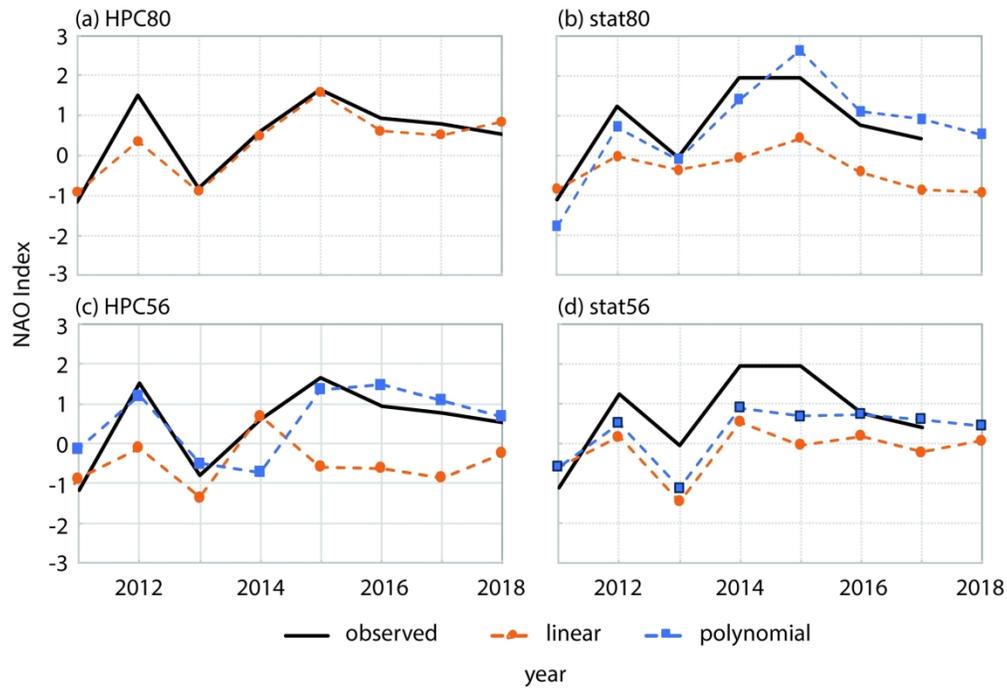


Figure 2. Out-of-sample forecasts compared with observations for linear (LIN) and polynomial (POLY) NARMAX models. Note there is no polynomial model for HPC80 and the observed station NAO is missing for 2018 in (b) and (d).

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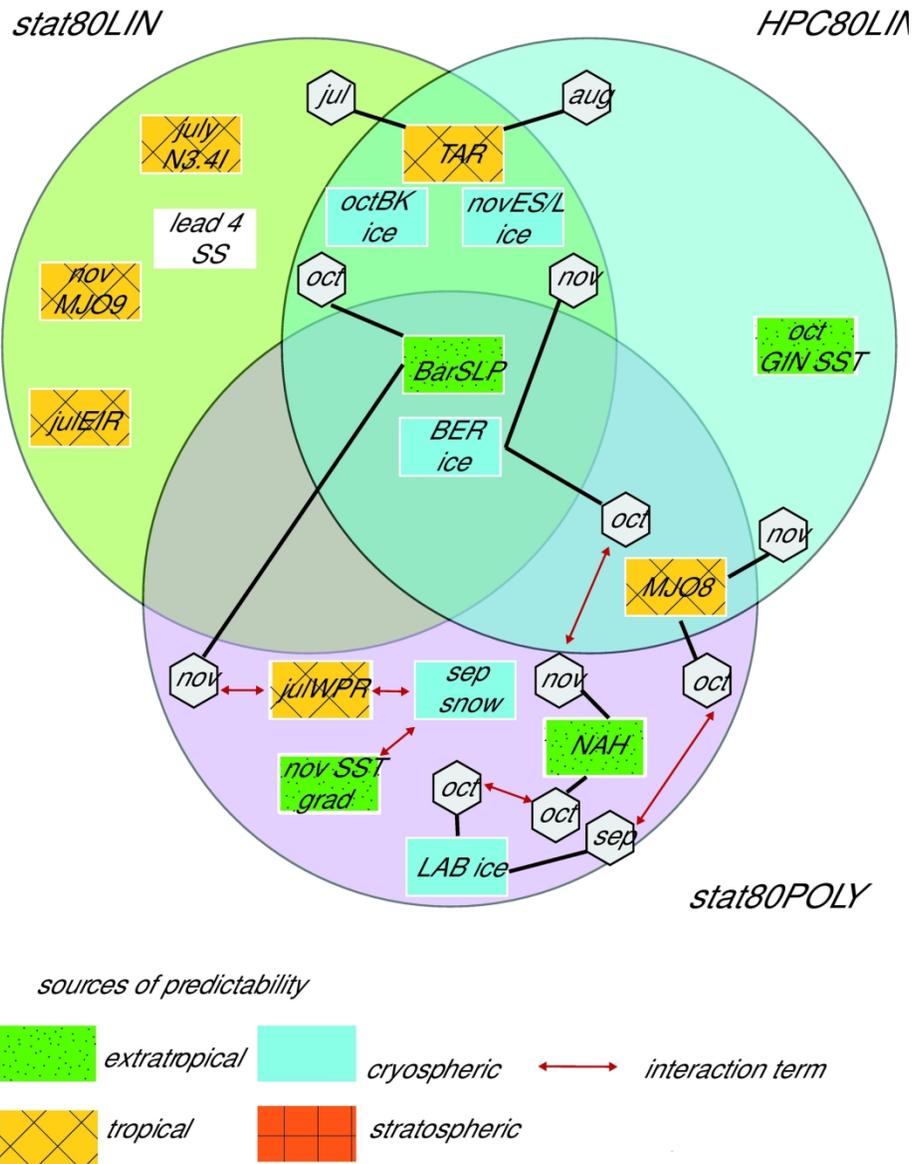


Figure 3. Predictors selected for the NAO models based on the 1980 training period. N3.4I =El Niño 3.4 discontinuous index; EIR=tropical East Indian Ocean rainfall, NAH=North Atlantic Horseshoe SST pattern; snow=Eurasian snowcover; BK ice=Barents-Kara Sea ice; GIN SST= Greenland-Iceland Norwegian SST; ES/L= East Siberian/Laptev Sea ice; TAR= tropical Atlantic rainfall; Bering ice=Bering Sea ice; MJO8/9=Phase 8 Madden-Julian Oscillation; Barents SLP=Barents Sea regional SLP; WPR=tropical West Pacific Ocean rainfall; SST gradient=North Atlantic SST gradient; LAB ice=Labrador Sea ice; lead 4 SS =sunspot cycle leading by 4 years. Where predictor month is not specified, it is indicated by a white hexagon linked to the variable. This is used where different models select different months of a common predictor.

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**Table I.** Summary of datasets used. For more detailed information, see Table S1.

dataset	Obtained from	Variable used
Hurrell PC_NAO	<a href="https://climatedataguide.ucar.edu/climate-data/hurrell-north-atlantic-oscillation-nao-index-pc-based">https://climatedataguide.ucar.edu/climate-data/hurrell-north-atlantic-oscillation-nao-index-pc-based</a>	DJF NAO index
Station-based NAO	<a href="https://rda.ucar.edu/datasets/ds570.0">https://rda.ucar.edu/datasets/ds570.0</a>	MSLP, Azores and Iceland
AMO	<a href="http://www.climexp.knmi.nl">www.climexp.knmi.nl</a>	HadSST3.1.1 SST
HadISST1	<a href="http://www.climexp.knmi.nl">www.climexp.knmi.nl</a>	SST, SIC
GPCPv2.3	<a href="http://www.climexp.knmi.nl">www.climexp.knmi.nl</a>	Tropical precipitation
Carbon dioxide	<a href="http://www.esrl.noaa.gov/gmd/ccgg/trends/data.html">www.esrl.noaa.gov/gmd/ccgg/trends/data.html</a>	Annual CO <sub>2</sub> level
QBO	<a href="http://www.geo.fu-berlin.de/en/met/ag/strat/produkte/qbo/index.html">www.geo.fu-berlin.de/en/met/ag/strat/produkte/qbo/index.html</a>	Mean zonal wind, 30hPa
sunspots	<a href="http://sidc.oma.be">http://sidc.oma.be</a>	Sunspot number
JRA-55 SPV_T100	Blanca Ayarzagüena, University of Exeter	Temperature 100hPa
NCEP/NCAR SLP	<a href="http://www.climexp.knmi.nl">www.climexp.knmi.nl</a>	Sea level pressure
Rutgers Global Snow Lab snow cover extent	<a href="https://climate.rutgers.edu/snowcover/docs.php?target=datareq">https://climate.rutgers.edu/snowcover/docs.php?target=datareq</a>	Snow cover extent
HadCRUT4.6	<a href="https://www.metoffice.gov.uk/hadobs/hadcrut4/">https://www.metoffice.gov.uk/hadobs/hadcrut4/</a>	2m temperature anomaly
MJO Indices	<a href="http://www.cpc.ncep.noaa.gov/products/precip/CWlink/daily_mjo_index/pentad.html">www.cpc.ncep.noaa.gov/products/precip/CWlink/daily_mjo_index/pentad.html</a>	200hPa velocity potential anomalies

<b>a) stat80LIN</b>		<b>Model Parameter</b>		
		<b>7-term</b>	<b>8-term</b>	<b>9-term</b>
	<b>Model Term</b>			
1	July N3.4I	-0.77	-0.39	-0.23
2	July tropical Atlantic rainfall	-0.10	-0.16	-0.22
3	November Bering Sea ice	0.40	0.46	0.55
4	October Barents-Kara Sea ice	0.53	0.67	0.73
5	October Barents Sea SLP	-0.30	-0.38	-0.37
6	November East Siberian-Laptev Sea ice	-0.32	-0.29	-0.34
7	November MJO phase 9	-0.38	-0.53	-0.63
8	July tropical East Indian Ocean rainfall		0.29	0.30
9	lead 4 year sunspot cycle			-0.18
<b>Individual Model Performance</b>				
Training	MAE	0.29	0.27	0.22
data	RMSE	0.37	0.31	0.27
	correlation	0.92	0.94	0.95
Testing data	MAE	1.02	1.25	1.07
	RMSE	1.14	1.38	1.26
	correlation	0.77	0.83	0.81

<b>b)HPC80LIN</b>		<b>Model Parameter</b>		
		<b>7-term</b>	<b>8-term</b>	<b>9-term</b>
	<b>Model Term</b>			
1	October Barents Kara Sea ice	0.91	0.94	0.85
2	October Barents Sea SLP	-0.42	-0.47	-0.48
3	November Bering Sea ice	0.57	0.35	0.32
4	November East Siberian-Laptev Sea ice	-0.44	-0.47	-0.44
5	constant	0.28	0.27	0.28
6	October GIN SST	0.33	0.42	0.39
7	November MJO phase 8	-0.46	-0.62	-0.57
8	October Bering Sea ice		0.35	0.37
9	August tropical Atlantic rainfall			0.23
<b>Individual Model Performance</b>				
Training	MAE	0.33	0.29	0.23
data	RMSE	0.41	0.35	0.28
	correlation	0.93	0.95	0.97
Testing data	MAE	0.33	0.30	0.36
	RMSE	0.49	0.47	0.48
	correlation	0.92	0.88	0.88

<b>c) stat56LIN</b>		<b>Model Parameter</b>		
	<b>Model Term</b>	<b>6-term</b>	<b>7-term</b>	<b>8-term</b>
1	May Greenland Sea ice	-0.15	-0.16	-0.14
2	September NAH	-0.27	-0.27	-0.27
3	July N3.4I	-0.61	-0.65	-0.61
4	October Barents-Kara Sea ice	0.37	0.41	0.44
5	October Barents Sea SLP	-0.23	-0.22	-0.25
6	November tropical West Indian Ocean SST	0.14	0.16	0.14
7	November East Siberian-Laptev Sea ice		-0.16	-0.15
8	October Greenland Sea ice			-0.08
<b>Individual Model Performance</b>				
Training data	MAE	0.52	0.51	0.51
	RMSE	0.66	0.65	0.64
	correlation	0.65	0.67	0.69
Testing data	MAE	1.27	0.95	1.03
	RMSE	1.38	1.07	1.14
	correlation	0.56	0.74	0.73

<b>d) HPC56LIN</b>		<b>Model Parameter</b>		
	<b>Model Term</b>	<b>5-term</b>	<b>6-term</b>	<b>7-term</b>
1	October NAH	-0.36	-0.33	-0.28
2	November SPV	-0.38	-0.29	-0.28
3	October SPV	0.41	0.34	0.36
4	July Greenland Sea ice	-0.12	-0.22	-0.27
5	October Barents Sea SLP	-0.31	-0.41	-0.41
6	October Barents-Kara Sea ice		0.44	0.44
7	November Bering Sea ice			0.18
<b>Individual Model Performance</b>				
Training data	MAE	0.68	0.61	0.58
	RMSE	0.79	0.72	0.69
	correlation	0.75	0.80	0.81
Testing data	MAE	0.83	1.22	1.29
	RMSE	0.99	1.44	1.56
	correlation	0.41	0.49	0.42

**Table II.** Selected predictors, model coefficients and verification statistics for linear NARMAX models.

a) models	MAE		RMSE		correlation		MSESS	D
	training	testing	training	testing	training	testing		
<b>1980 linear</b>								
stat80LIN	0.25	1.11	0.29	1.26	0.95	0.82	-0.96	0.80
HPC80LIN	0.27	0.29	0.31	0.48	0.96	0.90	0.82	0.75
<b>1980 polynomial</b>								
stat80POLY	0.35	<b>0.43</b>	0.41	<b>0.52</b>	0.89	<b>0.92</b>	<b>0.68</b>	<b>0.89</b>

b) models	MAE		RMSE		correlation		MSESS	D
	training	testing	training	testing	training	testing		
<b>1956 linear</b>								
stat56LIN	0.51	1.08	0.64	1.19	0.76	0.69	-0.44	0.80
HPC56LIN	0.61	1.10	0.71	1.31	0.81	0.46	-0.28	0.68
<b>1956 NARMAX</b>								
stat56POLY	<b>0.40</b>	<b>0.68</b>	<b>0.50</b>	<b>0.81</b>	<b>0.87</b>	<b>0.77</b>	<b>0.34</b>	0.79
HPC56POLY	<b>0.40</b>	<b>0.59</b>	<b>0.52</b>	<b>0.70</b>	<b>0.90</b>	<b>0.73</b>	<b>0.68</b>	<b>0.79</b>

**Table III.** Verification statistics for averaged linear and polynomial NARMAX models, for a) periods 1980-2018 and b) 1956-2018. Bold figures show if polynomial outperforms the linear model.

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<b>a)</b> <b>stat80POLY Model Term</b>		<b>Model Parameter</b>		
		<b>4-term</b>	<b>5-term</b>	<b>6-term</b>
1	Jul tropical West Pacific Rainfall* Nov Barents Sea SLP	-0.34	-0.31	-0.33
2	Oct Labrador Sea ice LAB * Oct NAH	0.63	0.64	0.58
3	Sep snow * Nov_SST gradient	0.48	0.55	0.56
4	Oct Bering Sea ice * Nov NAH	0.22	0.22	0.27
5	Sep Labrador Sea ice * Oct MJO phase 8		0.51	0.51
6	Jul tropical West Pacific Rainfall * Sep snow			0.33
<b>Individual Model Performance</b>				
Training	MAE	0.40	0.36	0.33
data	RMSE	0.48	0.43	0.38
	correlation	0.84	0.88	0.91
Testing data	MAE	0.43	0.53	0.48
	RMSE	0.50	0.58	0.62
	correlation	0.90	0.88	0.95

<b>b)</b> <b>stat56POLY Model Term</b>		<b>Parameter</b>		
		<b>8-term</b>	<b>9-term</b>	<b>10-term</b>
1	Jul tropical West Indian Ocean SST * Sep GIN SST	-0.06	-0.05	-0.07
2	Nov Bering Sea ice	-0.33	-0.67	-0.49
3	Sep NAH			
4	Nov East Siberian-Laptev Sea ice * Nov GIN SST	-0.28	-0.28	-0.27
5	Jul Greenland Sea ice * Jul Greenland Sea ice	-0.25	-0.19	-0.18
6	Oct Barents-Kara Sea ice	-0.03	-0.03	-0.03
7	Oct Barents Sea SLP * Nov GIN SST	0.37	0.33	0.32
8	Oct Barents Sea SLP			
9	Jul N3.4I * Sep Canadian Archipelago-Baffin sea ice	-0.21	-0.22	-0.24
10	Jul N3.4I * Sep NAH	-0.22	-0.20	-0.21
			0.19	0.18
				0.46
<b>Individual Model Performance</b>				
Training	RMSE	0.53	0.51	0.49
data	MAE	0.43	0.40	0.40
	Correlation	0.85	0.86	0.87
Testing	RMSE	0.74	0.85	0.85
data	MAE	0.67	0.68	0.69
	Correlation	0.70	0.63	0.82

<b>c)</b>		<b>Parameter</b>		
<b>HPC56POLY Model Term</b>		<b>6-term</b>	<b>7-term</b>	<b>8-term</b>
1	Oct Labrador Sea ice * Oct NAH	0.45	0.41	0.43
2	Aug N3.4 * lead 3 year sunspot cycle	0.63	0.69	0.73
3	atmospheric CO <sub>2</sub> * Jun Greenland Sea ice	0.04	0.04	0.03
4	Nov N3.4I * Sep Barents Sea SLP	0.82	0.97	1.19
5	Jun Hudson Bay sea ice * volcanic index	1.00	1.00	0.99
6	Sep GIN SST * Oct SST gradient	0.25	0.29	0.32
7	Sep QBO * Nov Barents Sea SLP		0.26	0.28
8	Sep N3.4I * Nov GIN SST			-0.56
<b>Individual Model Performance</b>				
Training	RMSE	0.59	0.55	0.49
data	MAE	0.45	0.42	0.37
	Correlation	0.86	0.89	0.91
Testing	RMSE	0.60	0.68	0.75
data	MAE	0.52	0.53	0.57
	Correlation	0.79	0.73	0.68

**Table IV.** As for Table II, but for polynomial models.

**Graphical Abstract****Complex Systems Modelling for Statistical Forecasting of Winter North Atlantic****Atmospheric Variability: a New Approach****Richard J. Hall\*, Hua-Liang Wei, Edward Hanna**

A new approach to seasonal forecasting based on complex systems modelling is presented.

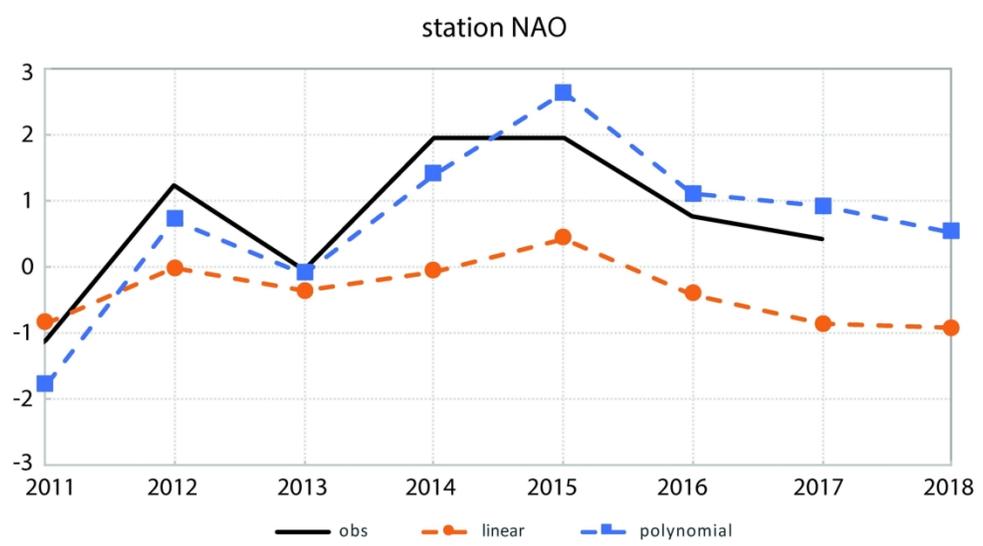
The focus is on North Atlantic winter atmospheric circulation, specifically the NAO.

Polynomial models show greater skill than linear versions and out-of-sample forecasts show promising skill, closely matching the observed time series. Potential non-linear interactions between predictors are identified.

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